<b>bitest</b> — Binomial probability test							
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# Description

bitest performs exact hypothesis tests for binomial random variables. The null hypothesis is that the probability of a success on a trial is  $\#_p$ . The total number of trials is the number of nonmissing values of *varname* (in bitest) or  $\#_N$  (in bitesti). The number of observed successes is the number of 1s in *varname* (in bitest) or  $\#_{succ}$  (in bitesti). *varname* must contain only 0s, 1s, and missing.

bitesti is the immediate form of bitest; see [U] 19 Immediate commands for a general introduction to immediate commands.

# **Quick start**

Exact test for probability of success (a = 1) is 0.4

```
bitest a = .4
```

With additional exact probabilities

```
bitest a = .4, detail
```

Exact test that the probability of success is 0.46, given 22 successes in 74 trials bitesti 74 22 .46

# Menu

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### bitesti

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## Syntax

Binomial probability test

bitest varname== #p [if] [in] [weight] [, detail]

Immediate form of binomial probability test

bitesti  $\#_N \#_{succ} \#_p [, \underline{d}etail]$ 

by and collect are allowed with bitest; see [U] **11.1.10 Prefix commands**. fweights are allowed with bitest; see [U] **11.1.6 weight**.

# Option

Advanced

detail shows the probability of the observed number of successes,  $k_{obs}$ ; the probability of the number of successes on the opposite tail of the distribution that is used to compute the two-sided *p*-value,  $k_{opp}$ ; and the probability of the point next to  $k_{opp}$ . This information can be safely ignored. See the technical note below for details.

## **Remarks and examples**

Remarks are presented under the following headings:

bitest bitesti

### bitest

#### Example 1

We test 15 university students for high levels of one measure of visual quickness which, from other evidence, we believe is present in 30% of the nonuniversity population. Included in our data is quick, taking on the values 1 ("success") or 0 ("failure") depending on the outcome of the test.

. use https://www.stata-press.com/data/r19/quick . bitest quick == 0.3 Binomial probability test Variable Ν Observed k Expected k Assumed p Observed p quick 15 7 4.5 0.30000 0.46667  $Pr(k \ge 7)$ = 0.131143 (one-sided test)  $Pr(k \le 7)$ = 0.949987(one-sided test)  $Pr(k \le 1 \text{ or } k \ge 7) = 0.166410$ (two-sided test)

The first part of the output reveals that, assuming a true probability of success of 0.3, the expected number of successes is 4.5 and that we observed seven. Said differently, the assumed frequency under the null hypothesis  $H_0$  is 0.3, and the observed frequency is 0.47.

The first line under the table is a one-sided test; it is the probability of observing seven or more successes conditional on p = 0.3. It is a test of  $H_0: p = 0.3$  versus the alternative hypothesis  $H_A: p > 0.3$ . Said in English, the alternative hypothesis is that more than 30% of university students score at high levels on this test of visual quickness. The *p*-value for this hypothesis test is 0.13.

The second line under the table is a one-sided test of  $H_0$  versus the opposite alternative hypothesis  $H_A$ : p < 0.3.

The third line is the two-sided test. It is a test of  $H_0$  versus the alternative hypothesis  $H_A: p \neq 0.3$ .

### Technical note

The *p*-value of a hypothesis test is the probability (calculated assuming  $H_0$  is true) of observing any outcome as extreme or more extreme than the observed outcome, with extreme meaning in the direction of the alternative hypothesis. In example 1, the outcomes k = 8, 9, ..., 15 are clearly "more extreme" than the observed outcome  $k_{obs} = 7$  when considering the alternative hypothesis  $H_A: p \neq 0.3$ . However, outcomes with only a few successes are also in the direction of this alternative hypothesis. For two-sided hypotheses, outcomes with k successes are considered "as extreme or more extreme" than the observed outcome  $k_{obs}$  if  $Pr(k) \leq Pr(k_{obs})$ . Here Pr(k = 0) and Pr(k = 1) are both less than Pr(k = 7), so they are included in the two-sided *p*-value.

The detail option allows you to see the probability (assuming that  $H_0$  is true) of the observed successes (k = 7) and the probability of the boundary point (k = 1) of the opposite tail used for the two-sided *p*-value.

. bitest quick == 0.3,	detail				
Binomial probability te	st				
Variable	Ν	Observed k	Expected 1	k Assumed p	Observed p
quick	15	7	4.	5 0.30000	0.46667
	= 0.94	1143 (one- 9987 (one- 6410 (two-	sided test)		
	= 0.08 = 0.09	1130 (obse 1560	rved)		
Pr(k == 1)	= 0.03	0520 (oppo	site extreme	)	

Also shown is the probability of the point next to the boundary point. This probability, namely, Pr(k = 2) = 0.092, is certainly close to the probability of the observed outcome Pr(k = 7) = 0.081, so some people might argue that k = 2 should be included in the two-sided *p*-value. Statisticians (at least some we know) would reply that the *p*-value is a precisely defined concept and that this is an arbitrary "fuzzification" of its definition. When you compute exact *p*-values according to the precise definition of a *p*-value, your type I error is never more than what you say it is—so no one can criticize you for being anticonservative. Including the point k = 2 is being overly conservative because it makes the *p*-value larger yet. But it is your choice; being overly conservative, at least in statistics, is always safe. Know that bitest and bitesti always keep to the precise definition of a *p*-value, so if you wish to include this extra point, you must do so by hand or by using the r() stored results; see *Stored results* below.

 $\triangleleft$ 

### bitesti

#### Example 2

The binomial test is a function of two statistics and one parameter: N, the number of observations;  $k_{obs}$ , the number of observed successes; and p, the assumed probability of a success on a trial. For instance, in a city of N = 2,500,000, we observe  $k_{obs} = 36$  cases of a particular disease when the population rate for the disease is p = 0.00001.

```
. bitesti 2500000 36 .00001
Binomial probability test
            Ν
                 Observed k
                               Expected k
                                             Assumed p
                                                          Observed p
    2,500,000
                         36
                                        25
                                               0.00001
                                                             0.00001
                                       (one-sided test)
 Pr(k >= 36)
                          = 0.022458
 Pr(k <= 36)
                          = 0.985448
                                       (one-sided test)
 Pr(k \le 14 \text{ or } k \ge 36) = 0.034859 (two-sided test)
```

### Stored results

bitest and bitesti store the following in r():

Scalars

r(N)	number $N$ of trials
r(P_p)	assumed probability $p$ of success
r(k)	observed number $k$ of successes
r(p_1)	lower one-sided p-value
r(p_u)	upper one-sided p-value
r(p)	two-sided p-value
r(k_opp)	opposite extreme k
r(P_k)	probability of observed $k$ (detail only)
r(P_oppk)	probability of opposite extreme $k$ (detail only)
r(k_nopp)	k next to opposite extreme (detail only)
r(P_noppk)	probability of k next to opposite extreme (detail only)

## Methods and formulas

Let N,  $k_{obs}$ , and p be, respectively, the number of observations, the observed number of successes, and the assumed probability of success on a trial. The expected number of successes is Np, and the observed probability of success on a trial is  $k_{obs}/N$ .

bitest and bitesti compute exact *p*-values based on the binomial distribution. The upper one-sided *p*-value is

$$\Pr(k \geq k_{\rm obs}) = \sum_{m=k_{\rm obs}}^{N} \binom{N}{m} p^m (1-p)^{N-m}$$

The lower one-sided *p*-value is

$$\Pr(k \leq k_{\rm obs}) = \sum_{m=0}^{k_{\rm obs}} \binom{N}{m} p^m (1-p)^{N-m}$$

If  $k_{obs} \ge Np$ , the two-sided *p*-value is

$$\Pr(k \le k_{opp} \text{ or } k \ge k_{obs})$$

where  $k_{opp}$  is the largest number  $\leq Np$  such that  $Pr(k = k_{opp}) \leq Pr(k = k_{obs})$ . If  $k_{obs} < Np$ , the two-sided *p*-value is

$$\Pr(k \le k_{obs} \text{ or } k \ge k_{opp})$$

where  $k_{opp}$  is the smallest number  $\geq Np$  such that  $\Pr(k = k_{opp}) \leq \Pr(k = k_{obs})$ .

## Reference

Hoel, P. G. 1984. Introduction to Mathematical Statistics. 5th ed. New York: Wiley.

## Also see

- [R] ci Confidence intervals for means, proportions, and variances
- [R] prtest Tests of proportions

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